

# A Genome-Wide Association Meta-Analysis of Attention-Deficit/Hyperactivity Disorder Symptoms in Population-Based Paediatric Cohorts

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## Abstract

**Objective**—To elucidate the influence of common genetic variants on childhood attention-deficit/hyperactivity disorder (ADHD) symptoms, to identify genetic variants that explain its high heritability, and to investigate the genetic overlap of ADHD symptom scores with ADHD diagnosis.

**Method**—Within the EARly Genetics and Lifecourse Epidemiology (EAGLE) consortium, genome-wide single nucleotide polymorphisms (SNPs) and ADHD symptom scores were available for 17,666 children (< 13 years) from nine population-based cohorts. SNP-based heritability was estimated in data from the three largest cohorts. Meta-analysis based on genome-wide association (GWA) analyses with SNPs was followed by gene-based association tests, and the overlap in results with a meta-analysis in the Psychiatric Genomics Consortium (PGC) case-control ADHD study was investigated.

**Results**—SNP-based heritability ranged from 5% to 34%, indicating that variation in common genetic variants influences ADHD symptom scores. The meta-analysis did not detect genome-wide significant SNPs, but three genes, lying close to each other with SNPs in high linkage disequilibrium (LD), showed a gene-wide significant association ( $p$  values between  $1.46 \times 10^{-6}$  and  $2.66 \times 10^{-6}$ ). One gene, *WASL*, is involved in neuronal development. Both SNP- and gene-based analyses indicated overlap with the PGC meta-analysis results with the genetic correlation estimated at 0.96.

**Conclusion**—The SNP-based heritability for ADHD symptom scores indicates a polygenic architecture and genes involved in neurite outgrowth are possibly involved. Continuous and dichotomous measures of ADHD appear to assess a genetically common phenotype. A next step is to combine data from population-based and case-control cohorts in genetic association studies to increase sample size and improve statistical power for identifying genetic variants.

## Keywords

GWA; SNP heritability; attention problems; ADHD symptoms; meta-analysis

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## Introduction

Attention-deficit/hyperactivity disorder (ADHD) is a common psychiatric condition in childhood with a prevalence of around five percent across countries worldwide.<sup>1</sup> As an objective diagnostic test is lacking, diagnoses are based on the occurrence of age-

inappropriate impulsive, hyperactive, and inattentive behaviors that occur in multiple settings and cause significant impairment.<sup>2,3</sup> It is well established that genetic factors explain a large part of the individual differences in the vulnerability for ADHD. The heritability of childhood ADHD and related traits, such as continuous measures of attention problems and hyperactivity, has been estimated at around 75% (see e.g. Kan et al. 2013<sup>4</sup> and Lichtenstein et al. 2010<sup>4</sup>). Consequently, several studies now aim to identify genetic variants for ADHD. Ten candidate genes show replicated evidence for association according to a recent review.<sup>5</sup> In genome-wide association (GWA) studies, the test of the effects of single genetic variants has not yet yielded genome-wide significant hits, but a gene-enrichment analysis, including single nucleotide polymorphisms (SNPs) showing genome-wide suggestive signals, pointed to several biological pathways involved in neural processes, such as neurodevelopment.<sup>5</sup> Additional evidence for the role of common SNPs (SNPs with a frequency above 5%) in ADHD comes from polygenic analyses in which the joint effect of a large number of SNPs or all SNPs is estimated. All studies but one found that the SNPs explained a significant proportion of the variance,<sup>6–13</sup> suggesting that associated genetic variants are likely to be detected in larger GWA meta-analyses. This is confirmed by the results from the latest meta-analysis of the Psychiatric Genomics Consortium (PGC) ADHD subgroup that yielded several genome-wide significant hits (D Demontis for the PGC ADHD subgroup: presentation 23rd World Congress of Psychiatric Genetics, October 2015, Toronto, CA).

Another recommendation for future gene-finding studies has been to incorporate dimensional approaches of ADHD, such as continuous measures of ADHD symptoms.<sup>5</sup> This is supported by polygenic risk score analyses showing that individuals' polygenic risk scores based on the effects of SNPs in a GWA analysis in ADHD case-control studies significantly predicted continuous ADHD symptom scores, and vice versa.<sup>13–15</sup> Other studies have suggested that a diagnosis of ADHD can be regarded as the extreme end of a continuous distribution of inattentive and hyperactive behaviors,<sup>16–18</sup> and twin studies also showed a substantial overlap between the genetic factors for a clinical diagnosis of ADHD and continuous measures of ADHD symptoms in the general population.<sup>17,18</sup>

Many population-based paediatric cohorts have collected genome-wide SNP data and continuous ADHD symptom scores, providing an underused opportunity for gene-finding studies for ADHD. Case-control studies benefit from oversampling the high-scoring end of the distribution, but analyzing the full information on symptom severity in the population can also be a powerful approach, especially for a relatively common disorder as ADHD.<sup>13,19</sup> The current report describes the first GWA meta-analysis of continuous measures of ADHD in 17,666 children from nine population-based cohorts. To investigate the polygenic nature of the phenotype, we estimated the variance in attention problems that was explained by all SNPs. Next, SNP and gene-based association analyses were performed. We specifically looked at whether the ten candidate genes identified in the recent review showed evidence for association.<sup>5</sup> Finally, to examine the overlap in genetic influences on ADHD diagnosis and ADHD symptom scores, we investigated the concordance in the effects of SNPs and genes as found in the current meta-analysis with the results of a meta-analysis based on case-control ADHD GWA studies, and estimated the genetic correlation.

## Method

### Cohorts

The EARly Genetics and Life course Epidemiology (EAGLE) consortium is a collaboration among several population-based birth cohorts from Europe, Australia, and the United States (<http://www.wikigenes.org/e/art/e/348.html>). The consortium focuses on a wide range of phenotypes in childhood.<sup>20,21</sup> EAGLE cohorts with ADHD symptom scores in childhood (age at measurement <13 years) were invited to participate in the meta-analysis. An overview of the nine cohorts included in the meta-analysis is provided in Table 1. Information on the individual cohorts can be found on their websites and in the publications listed in Table 1.

### Phenotype

Different instruments were used across cohorts (Table 1), including the Attention Problems scale of the Child Behavior Checklist (CBCL) and the Teacher Report Form (TRF), the Hyperactivity scale of the Strengths and Difficulties Questionnaire (SDQ), and the *DSM-IV* ADHD items as, for example, included in the Conners' Rating Scale (see Table S1, available online, for the items included in each scale).<sup>3,22–25</sup> For the meta-analysis, one phenotype was selected from each cohort. Based on the phenotype that was most available, school-age ratings were chosen over preschool-age ratings, parent ratings over teacher ratings, and the measurement instrument with the largest information density was preferred over the other instruments (Conners' *DSM-IV* > CBCL > SDQ).

### SNP-Based Heritability

The variance in ADHD symptom scores accounted for by the SNPs was estimated using Genomic-Relationship-Matrix Restricted Maximum Likelihood (GREML) as implemented in the Genomic Complex Trait Analysis (GCTA) software.<sup>26,27</sup> GREML is a linear mixed model that includes a genetic relatedness matrix (GRM) that contains a measure of genetic similarity between all possible pairs of (unrelated) individuals in a study. Genetic similarity is based on resemblance in SNP variants, hence the variance explained by the genetic relatedness matrix is often called the SNP heritability. Typically, only unrelated individuals (genetic relatedness < 0.025) are included in the construction of a GRM to prevent the estimate of the SNP heritability to be biased upwards.

GRM-based analyses were performed for the hyperactivity scale of the SDQ as measured in the Avon Longitudinal Study of Parents and Children (ALSPAC) at preschool (N=5,510) and at school-age (N=5,303), and for the Attention Problems scale of the CBCL 1.5-5 (N = 2,958) and the TRF (N=1,901) measured in the Netherlands Twin Register (NTR) and Generation R cohorts. These analyses were not performed in the other cohorts because of the smaller sample sizes.

In ALSPAC, the GRM was constructed based on observed genotypes. Sex, age at measurement, and two principal components were included as fixed effects in the model. In NTR, the CBCL 1.5-5 was assessed when the children were three years and in Generation R when they were six years of age. The NTR and Generation R samples were combined to

estimate the GRM. Individual-level genotype data from the NTR and Generation R were imputed together based on the Genome of the Netherlands reference set.<sup>28,29</sup> Sex, age at measurement, sample, and five principal components were included as fixed effects in the model.

### Data Quality Control and SNP and Gene-Based Association Meta-Analyses

Cohorts performed quality control (QC) and imputed their SNP genotype data using the March 2012 release of the 1000 Genomes reference set that includes all ethnicities.<sup>30</sup> Each cohort ran their own optimal pre-imputation genotype QC. An overview of the pre-imputation QC metrics and imputation methods applied in each cohort is provided in Table S2, available online. Briefly, filtering on sample and SNP call rate was similar between cohorts, with the exception of the Twins Early Development Study (TEDS), which had a lower threshold. The thresholds for Hardy-Weinberg Equilibrium, heterozygosity filtering, and other QC steps varied somewhat more. This likely has slightly decreased the final imputation quality of each cohort, as prior SNP filtering does not improve the imputation quality, since losing genotyped SNPs only makes the imputation worse.<sup>31</sup> Imputation with the Mach software has been shown to perform slightly better,<sup>31</sup> but differences tend to be small. All filtering decisions will eventually result in fewer SNPs in the meta-analysis, bringing the risk of a loss of potential signal but not leading to false-positive results.

A linear regression of the phenotype on sex, age at measurement, genotype dose, and principal components was performed in all cohorts. All cohorts analyzed data from unrelated individuals, except for the Netherlands Twin Register (NTR), which included both twins from a dizygotic twin pair and corrected standard errors in PLINK using the --family option (<http://pngu.mgh.harvard.edu/purcell/plink/>).<sup>32</sup> Table S2, available online, lists the analytic tools applied by each cohort.

Results were checked and meta-analyzed by two independent analysts. QC included calculation of the inflation factor lambda (the ratio of the observed versus the expected median chi-square), format checking, visual inspection of QQ plots, Manhattan plots, histograms of minor allele frequency (MAF) and INFO scores, consistency of reported allele frequency with the reference set, consistency of reported *p* value with reported beta and standard error (SE), and consistency of reported SE with reported sample size, standard deviation (SD), and MAF. All files were filtered using the software EasyQC<sup>33</sup> ([www.genepi-regensburg.de/easyqc](http://www.genepi-regensburg.de/easyqc)) based on  $R^2$  metric > 0.7 for MACH-based imputations and INFO metric > 0.8 for IMPUTE-based imputations. This filter was applied to all SNPs to ensure conservatism. In addition, SNPs were filtered for MAF > 0.03, Hardy Weinberg equilibrium  $p$  value < .0001, consistency of reported alleles and allele frequency with the reference set (maximum difference 0.2 with 1000G phase 1 v3), and duplicates (both occurrences removed).

As different phenotyping instruments were used across cohorts, the meta-analysis was based on *p* values and performed in METAL software (<http://www.sph.umich.edu/csg/abecasis/metal/>; option SCHEME SAMPLESIZE) including an application of genomic control to the results of the individual cohorts.<sup>34</sup> Meta-analysis results were filtered on a total sample size > 10,000. A *p* value <  $5 \times 10^{-8}$  was considered genome-wide significant.

Gene-based analyses were performed in MAGMA (Multi-marker Analysis of GenoMic Annotation).<sup>35</sup> In MAGMA, a gene test-statistic is calculated as the mean of the chi-square statistics for all the SNPs between the transcription start and stop sites of a gene. The gene  $p$  value is then obtained by using a known approximation of the sampling distribution. MAGMA corrects for gene size, number of SNPs in a gene, and linkage disequilibrium (LD) between SNPs in a gene using the SNP correlation matrix. We used the European ancestry samples from the 1000 Genomes project as reference data to estimate LD. Association was tested for 17,155 genes. The  $p$  value for gene-wide significance after Bonferroni correction was  $0.05/17,155=2.91 \times 10^{-6}$ .

### Overlap in Results Between the EAGLE and a PGC ADHD Case-Control Meta-Analysis

We investigated the overlap in the results for the SNP and gene-based analyses obtained in EAGLE with the results of the PGC ADHD meta-analysis (P Holmans for the PGC: presentation 21st World Congress of Psychiatric Genetics, October 2013, Boston, USA, <http://2013.ispg.net/wp-content/uploads/2013/10/Oral-Presentations-Abstract-Book.pdf>). The PGC sample comprised 5,621 cases and 13,589 controls. The overlap in SNP effects was investigated with SNP effect Concordance Analysis (SECA)<sup>36</sup> and with Linkage Disequilibrium Score (LDSC) Regression analysis.<sup>37,38</sup> Both methods only require the summary statistics of the association (meta)-analyses.

SECA takes the overlapping SNPs of both datasets and selects from each set of results SNPs with  $p$  values .01, .05, .1, .2, .3, .4, .5, .6, .7, .8, .9, and 1.0. This leads to a 12x12 matrix indicating the overlap in SNP effects for each combination of SNP sets in the two analyses. SECA performs several tests based on these 144 cells. We report the results of the Fisher exact tests analyzing for each combination of the SNP sets whether the number of SNPs that are concordant in the direction of effects for the two phenotypes is above chance. Next, an empirical  $p$  value is calculated that indicates whether the overlap is higher than expected by chance given the multiple testing (144 tests).

After testing for SNP concordance by SECA, we also tested for overlap at the gene level, following the procedure described in Zhao et al,<sup>39</sup> where gene-based associations were tested in GATES.<sup>40</sup> Independent genes were then identified by examining the LD between the most associated SNPs within each gene in the Genetic type I Error Calculator (GEC).<sup>41</sup> Exact binomial statistical tests then determined whether the number of genes with  $p$  values  $<.01$ ,  $<.05$  or  $<.1$  that were observed in both sets of results was significantly higher than expected.

The LDSC regression analysis was performed in the LDSC package (<https://github.com/bulik/ldsc>).<sup>37,38</sup> This analysis yields SNP heritability estimates of the traits and a genetic correlation between the traits.

## Results

### SNP-Based Heritability

The estimates of the SNP-based heritability are shown in Table 2. The estimates for the maternal ratings were 5% (non-significant [ns]) for preschool SDQ and 13% (ns) and 14%

( $p < .05$ ) for preschool CBCL and school-age SDQ, respectively. For teacher ratings, an SNP-based heritability of 34% ( $p < .05$ ) was observed. These results indicated that SNPs tag variants associated with various ADHD symptom scores.

### SNP and Gene-Based Meta-Analyses

The numbers of SNPs from each cohort that were included in the meta-analysis after QC are displayed in Table S3, available online. The QQ plot in Figure 1 shows the distribution of SNP  $p$  values from the meta-analysis filtered on SNPs that were present in at least 10,000 individuals. The lambda statistic of the meta-analysis was 0.98. Individual cohorts had lambdas  $\approx 1.08$  (Table S3, available online), implying absence of population stratification. The Manhattan plot in Figure 2 shows that none of the SNPs reached genome-wide significance. However, the QQ plot shows some departure from the expected line for the smallest  $p$  values, which may reflect the polygenic nature of the trait, i.e., many variants of small effects influencing ADHD symptoms. The strongest association was with rs56159542 on chromosome 19 ( $p = 1.48 \times 10^{-7}$ ). A summary of the top-signals that crossed the threshold of suggestive association at  $p < 1 \times 10^{-5}$  is included in Table 3a. Eight out of the top nine variants were located in genes. As shown in the locus zoom plots<sup>42</sup> of the top SNPs (Figure S1, available online), the signals of rs79846815, rs61227778, rs77216358 were restricted to one to three SNPs, suggesting that they might not be genuine signals.

Table 3b shows the top 10 genes observed in the gene-based analysis as performed in MAGMA. Three genes, *WASL*, *LMOD2*, and *ASB15*, yielded gene-wide associations. An SNP in *LMOD2* also yielded a suggestive association with ADHD symptom scores (Table 3a). The locus zoom plot (Figure S2, available online) shows that several SNPs show similar signals in the three genes as they are in a region with SNPs in high LD. Thus, the signals of the three genes are not independent of each other. The previously identified ADHD candidate genes did not show evidence for association.  $P$  values ranged from 0.11 to 0.91 (Table S4, available online).

### Overlap in Results Between the EAGLE and PGC ADHD Case-Control Meta-Analysis

The SECA software tested for 144 combinations of SNP subsets obtained from the EAGLE and the PGC ADHD meta-analyses, whether there were more SNPs showing concordance in the direction of effects than expected by chance (Fisher's Tests with OR = 1 and  $p$  value = .05). These analyses resulted in 111 SNP subsets with  $p < .05$ . This is significantly higher than expected by chance (empirical  $p$  value = .001), clearly indicating an overlap between the SNPs associated with ADHD symptom scores and ADHD diagnoses. This was confirmed by the analysis investigating the overlap in genes rather than SNPs.<sup>39</sup> The overlap in genes with a  $p$  value  $< .05$  or .1 in both cohorts was larger than expected by chance, with binomial  $p$  values of .05 and  $1.3 \times 10^{-3}$ , respectively.

In the LDSC regression analysis, the genetic correlation between the EAGLE ADHD symptom scores and the PGC ADHD case-control phenotype was estimated at 0.96 (se=0.28,  $p < .001$ ), with the SNP heritability for the EAGLE ADHD symptom scores estimated at 0.08 (se=0.03,  $p < .01$ ).

## Discussion

The current study comprised the largest GWA analysis of continuous ADHD symptom scores in children to date. We found that common variants included in GWA studies explained variation in ADHD symptom scores assessed in the general population. The SNP heritability estimates for the various measures from the participating cohorts ranged from 0.05 to 0.34. The SNP heritability based on the results of the meta-analysis in all cohorts with a total of 17,666 children was estimated at 8%. We did not detect SNPs at genome-wide significance levels, but detected three gene-wide significant results in the gene-based analyses. The analyses investigating the overlap in genetic influences on ADHD symptoms scores and ADHD diagnosis provided evidence for a considerable common genetic background with an estimate of the genetic correlation of 0.96.

The three associated genes, *LMOD2* (7q31.32), *ASB15* (7q31.32), and *WASL* (7q31.32), lie in a region with high LD, thereby making it difficult to decide on statistical grounds which gene contributes to this signal. Leiomodin (*LMOD*) is an actin-binding protein that acts as a filament nucleator in muscle cells.<sup>43</sup> Ankyrin repeat and SOCS box containing 15 (*ASB15*) gene product belongs to the ASB family of proteins that are part of a ubiquitination-mediated pathway.<sup>44</sup> The ubiquitin proteasome pathway has also been suggested to play a role in adult ADHD.<sup>45</sup> The protein encoded by Wiskott-Aldrich syndrome like (*WASL*) is involved in cytoskeletal organization during neuronal development, including long spine formation and neurite extension.<sup>46</sup> Given the enrichment of genes involved in directed neurite outgrowth in the analysis of the combined results of published GWA studies on ADHD, *WASL* seems the most likely candidate to drive the signal.

A power analysis in Quanto<sup>48</sup> suggested that the current sample has 80% power to detect a genome-wide significant effect explaining 0.21% of the variance, assuming an additive genetic effect, considering the largest possible sample of  $N=17,666$ , and assuming that a meta-analysis has as much statistical power as a single analysis of a similar sample size. The lack of genome-wide significant effects in combination with the observed SNP-based heritability estimates indicates that ADHD is likely to be highly polygenic, i.e., influenced by many common genetic variants of small effect sizes. These results are in line with earlier studies on ADHD based on case-control samples. Following Yang et al.,<sup>13</sup> we estimated that the PGC ADHD subgroup GWA meta-analysis of 5,621 cases and 13,589 controls had 8% more power than the current study, and yet no significant results were found despite the SNP heritability estimated at 28% in this sample.<sup>11</sup> Several other studies also found evidence for polygenicity of ADHD,<sup>6–9,49</sup> with the exception of one study.<sup>10</sup> The range of the SNP heritability estimates for the symptom scores in the current study (Table 2) seems quite large. However, these differences could well be due to chance judging by the large standard errors.

The provided evidence for the genetic overlap in continuous and dichotomous measures of ADHD agrees with findings from previous studies with smaller samples.<sup>13–15</sup> One of these studies also reported that an aggregate polygenic risk score derived from a sample of clinical cases of ADHD predicted preschool parent and school-age parent and teacher ratings of attention problems in a population-based cohort.<sup>49</sup> This indicates common genetic variance

across these measures and ages, which is confirmed by the SNP-based heritability of 8% that we calculated based on the results of the meta-analysis. Despite the use of various measures across the cohorts and the accompanying heterogeneity, there remains a signal after combining the results. Overall, it can be concluded that the different instruments assess an underlying common liability for ADHD. Therefore, combining various continuous ADHD measures assessed in the general population with dichotomous diagnosis of ADHD assessed in clinical samples can be a successful way to increase sample size and statistical power for GWA studies. This is supported by preliminary results of the PGC ADHD subgroup (R Walters: presentation 23rd World Congress of Psychiatric Genetics, October 2015, Toronto, CA).

Efforts to decrease heterogeneity across studies by harmonizing phenotypes can result in a further increase in power to detect genetic effects. Behavioral genetic studies have reported that genetic factors are not entirely similar across instruments, raters, and ages.<sup>50–54</sup> We ran additional SECA analyses to investigate the overlap in results of the current meta-analysis with the results from a GWAS in an independent sample of 727 Australian adolescents whose mothers provided retrospective ratings of their childhood attention skills and problems using the Strengths and Weaknesses of ADHD Symptoms and Normal Behavior Rating Scale (SWAN). These analyses did not show concordance in SNP effects, but we note this could be due to the small size of the Australian sample. Statistical methods like Item Response Theory (IRT) can be used to synchronize different measurement instruments in a sophisticated manner, and have already been successfully applied in a GWA meta-analysis of personality measures.<sup>55–57</sup> Another way to refine the phenotype when longitudinal data are available is to test the effect of the SNP on a latent variable that reflects stability over time and is more heritable than a measure at a single time point.<sup>58</sup>

To conclude, our results support the notion that ADHD is influenced by genes involved in neuronal development. By performing GWA meta-analyses in larger samples, we should be able to identify genetic variants for ADHD, further elucidating its biological foundation. The use of continuous ADHD symptom scores available in population-based cohorts is an exciting possibility to achieve this goal.

## Supplementary Material

Refer to Web version on PubMed Central for supplementary material.

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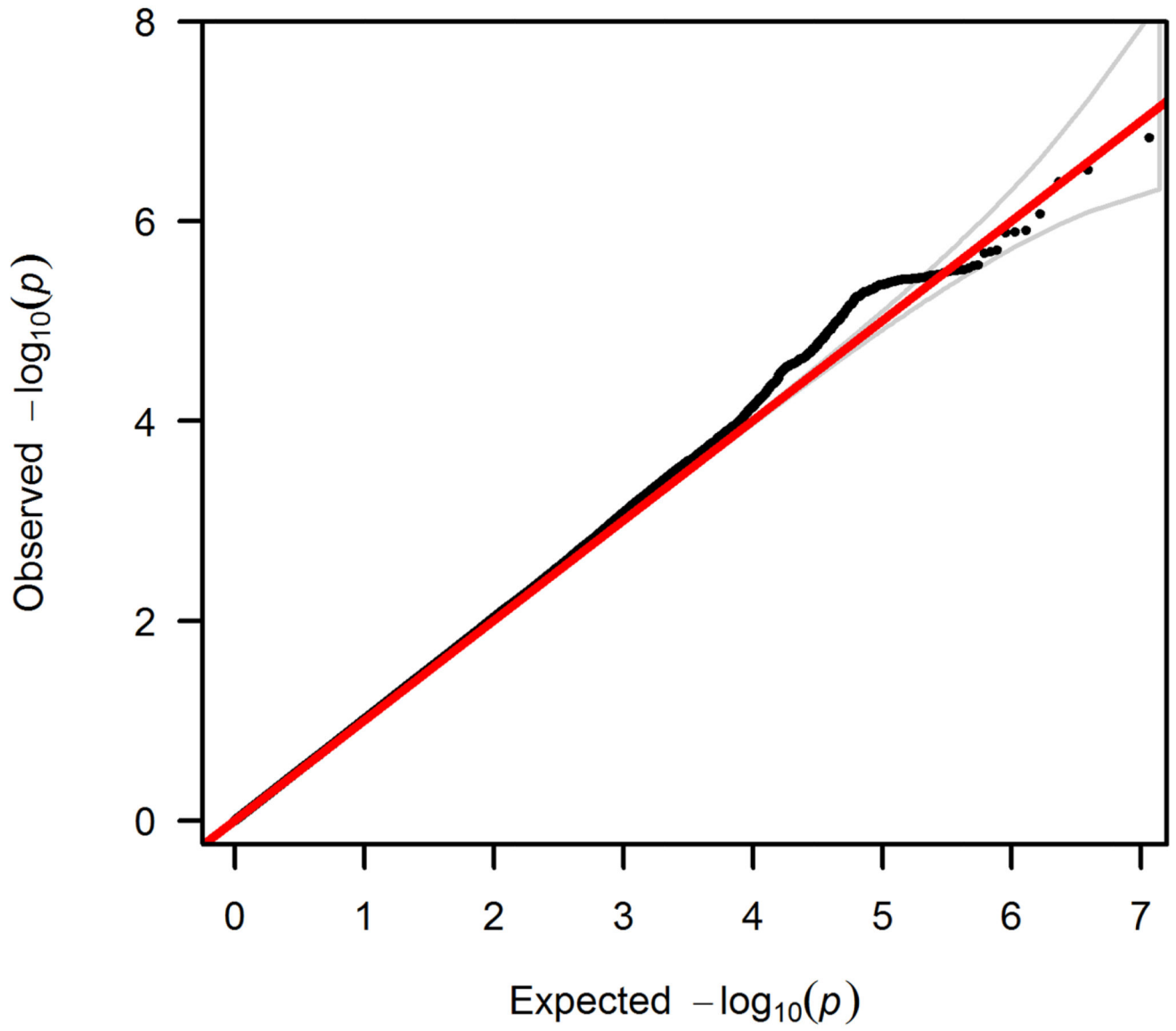
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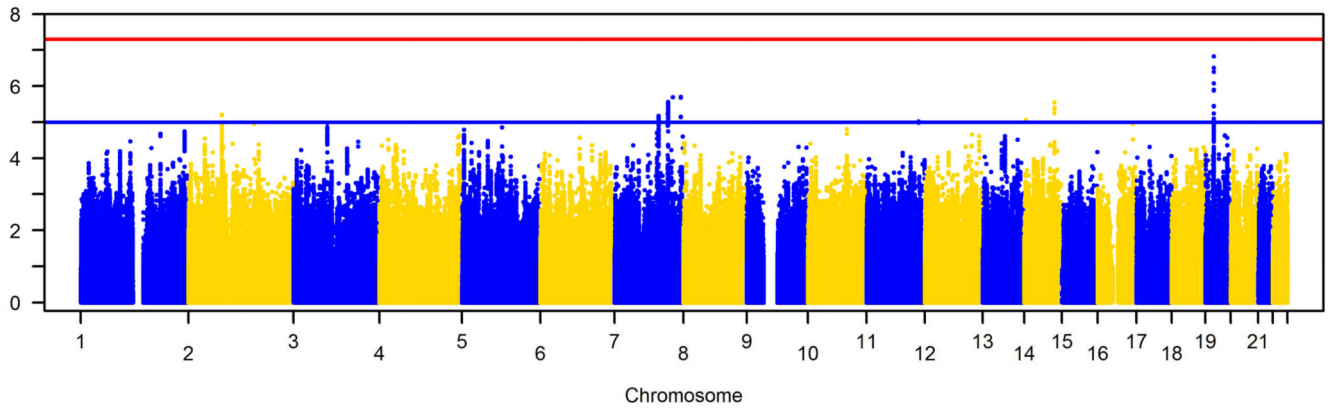
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**Figure 1.**  
QQ-plot of all meta-analysis results based on at least 10,000 individuals.





**Figure 2.** Manhattan plot of the meta-analysis results based on at least 10,000 individuals.

**Table 1**  
Description of the Cohorts and Attention-Deficit/Hyperactivity Disorder (ADHD) Instruments Included in the Meta-Analysis

Cohort	N	Phenotype instrument	Rater	Age in years mean (SD)	Sum score mean (SD)	Website	Reference papers
ALSPAC	5,757	SDQ	Parent	9.65 (0.12)	2.91 (2.24)	<a href="http://www.bristol.ac.uk/alspac/">www.bristol.ac.uk/alspac/</a>	59,60
Generation R	2,211	CBCL/1.5-5	Parent	6.01 (0.38)	1.38 (1.69)	<a href="http://www.generationr.nl">www.generationr.nl</a>	61
GINI / LISA	1,389	SDQ	Parent	10.04 (0.20)	2.71 (2.36)	<a href="http://www.helmholtz-muenchen.de/epi/arbeitsgruppen/umweltepidemiologie/projects-projekte/lisa-plus/index.html">http://www.helmholtz-muenchen.de/epi/arbeitsgruppen/umweltepidemiologie/projects-projekte/lisa-plus/index.html</a>	62,63
INMA	804	DSM-based scale	Teacher	4.91 (0.69)	5.38 (6.83)	<a href="http://www.proyectoinma.org/">www.proyectoinma.org/</a>	64
MoBa	665	CBCL/1.5-5	Parent	3.05 (0.10)	2.05 (1.67)	<a href="http://www.fhi.no/morogbarn">www.fhi.no/morogbarn</a>	65
NTR	1,605	CBCL/6-12	Parent	9.95 (0.85)	3.25 (3.39)	<a href="http://www.tweelingenregister.org">www.tweelingenregister.org</a>	66
Raine	1,344	CBCL/6-12	Parent	10.58 (0.20)	2.60 (3.17)	<a href="http://www.rainestudy.org.au">www.rainestudy.org.au</a>	67-69
TEDS	2,606	Conners'	Parent	7.88 (0.52)	10.51 (8.62)	<a href="http://www.teds.ac.uk">www.teds.ac.uk</a>	70
TRAILS	1,285	CBCL/6-12	Parent	11.08 (0.56)	4.27 (3.40)	<a href="http://www.trails.nl">www.trails.nl</a>	71

Note: ALSPAC = Avon Longitudinal Study of Parents and Children; CBCL = Child Behavior Check List; GINI = German Infant Nutritional Intervention; INMA = Infancia y Medio Ambiente; LISA = Influence of Life-style factors on Development of the Immune System and Allergies in East and West Germany plus Air Pollution and Genetics on Allergy Development; MoBa = Norwegian Mother and Child Cohort Study; NTR = Netherlands Twin Register; SDQ = Strengths and Difficulties Questionnaire; TEDS = Twins Early Development Study; TRAILS = 'Tracking Adolescents' Individual Lives' Survey.

**Table 2**

Genomic Complex Trait Analysis Single Nucleotide Polymorphisms Heritabilities for Mother<sup>a</sup>- and Teacher<sup>b</sup>-Rated Attention-Deficit/Hyperactivity Disorder (ADHD) Symptoms in the Avon Longitudinal Study of Parents and Children (ALSPAC) and the Combined Generation R and Netherlands Twin Register (NTR) Data

	Generation R/NTR		ALSPAC	
	CBCL (3 and 6 yrs)	TRF (7yrs)	SDQ 4 yrs	SDQ 9 yrs
n	2958	1901	5510	5303
Variance explained	0.13	0.34	0.05	0.14
SE	0.11	0.17	0.06	0.07
<i>p</i> -value	0.11	0.02	0.20	0.02

Note: CBCL = Child Behavior Checklist; SDQ = Strengths and Difficulties Questionnaire; TRF = Teacher Report Form.

<sup>a</sup>Preschool CBCL, preschool and school-age SDQ.

<sup>b</sup>TRF.

**Table 3a**  
**Top Signals From the Single Nucleotide Polymorphism (SNP) and Gene-Based Meta-Analyses**

3a. List of Independent Signals With  $p < 1 \times 10^{-5}$

SNP	Chr	Position (GRCh37)	Effect/Other Allele	Freq (freq in refset)	Total N	Direction of Effect	z-Score	p-value	Location in/to Nearest Gene
rs56159542	19	19682971	T/C	0.21 (0.19)	17666	- (-----)	-5.26	$1.48 \times 10^{-7}$	<i>PBX4</i> intronic
rs4629772	7	152823816	A/G	0.93 (0.93)	16322	- (----?--)	-4.76	$1.97 \times 10^{-6}$	downstream <i>ACTR3B</i>
rs79846815	7	134563570	A/T	0.97 (0.96)	11175	+ (+?+???)	4.75	$2.03 \times 10^{-6}$	<i>CALDI</i> intronic
rs7809453	7	123301940	A/G	0.54 (0.56)	17666	- (-----)	-4.69	$2.78 \times 10^{-6}$	<i>LMOD2</i> exonic, synonymous
rs79162905	14	89796072	A/G	0.11 (0.10)	17666	- (-----)	-4.68	$2.81 \times 10^{-6}$	<i>FOXN3</i> intronic
rs146855089	2	77317636	A/G	0.26 (0.27)	17666	- (-----)	-4.52	$6.18 \times 10^{-6}$	<i>LRRTM4</i> intronic
rs10808119	7	101840716	A/G	0.46 (0.45)	17666	+ (+++++)	4.50	$6.72 \times 10^{-6}$	<i>CUX1</i> intronic
rs61227778	14	24578916	A/G	0.95 (0.96)	10826	+ (+++++???)	4.45	$8.62 \times 10^{-6}$	<i>NRL</i> intronic
rs77216358	11	120311157	A/G	0.96 (0.97)	11042	- (--+???)	-4.43	$9.48 \times 10^{-6}$	<i>ARHGGEF12</i> intronic

**Table 3b**  
**Top Signals From the Single Nucleotide Polymorphism (SNP) and Gene-Based Meta-Analyses**

3b. The Top 10 Genes From the Gene-Based Tests in Multi-Marker Analysis of GenoMic Annotation (MAGMA)

Gene	Chr	Start Position (GRCh37)	Stop Position (GRCh37)	N SNPs	p-value
<b><i>LMOD2</i></b>	<b>7</b>	<b>123295861</b>	<b>123304147</b>	<b>13</b>	<b><math>1.46 \times 10^{-6}</math></b>
<b><i>WASL</i></b>	<b>7</b>	<b>123321997</b>	<b>123389116</b>	<b>117</b>	<b><math>1.50 \times 10^{-6}</math></b>
<b><i>ASB15</i></b>	<b>7</b>	<b>123249112</b>	<b>123277932</b>	<b>45</b>	<b><math>2.66 \times 10^{-6}</math></b>
<i>CUX1</i>	7	101459184	101927250	1003	$6.03 \times 10^{-5}$
<i>HAPLN4</i>	19	19366450	19373596	13	$6.10 \times 10^{-5}$
<i>CILP2</i>	19	19649074	19657468	16	$6.35 \times 10^{-5}$
<i>LRRTM4</i>	2	76974849	77749502	1858	$8.50 \times 10^{-5}$
<i>ZNF234</i>	19	44645710	44664462	19	$8.73 \times 10^{-5}$
<i>NDUFA13</i>	19	19627019	19639013	19	$1.09 \times 10^{-4}$
<i>RWDD4</i>	4	184560789	184580331	41	$1.52 \times 10^{-4}$

Note: Boldface text is gene-wide significant ( $p < 2.91 \times 10^{-6}$ ).